TWO GRUMPY GIANTS AND A BABY

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Abstract. Pollard's rho algorithm, along with parallelized, vectorized, and negating variants, is the standard method to compute discrete logarithms in generic prime-order groups. This paper presents two reasons that Pollard's rho algorithm is farther from optimality than generally believed. First, "higher-degree local anti-collisions" make the rho walk less random than the predictions made by the conventional Brent–Pollard heuristic. Second, even a truly random walk is suboptimal, because it suffers from "global anti-collisions" that can at least partially be avoided. For example, after $(1.5 + o(1))\sqrt{\ell}$ additions in a group of order ℓ (without fast negation), the baby-step-giant-step method has probability $0.5625 + o(1)$ of finding a uniform random discrete logarithm; a truly random walk would have probability $0.6753\ldots + o(1)$; and this paper's new two-grumpy-giants-and-a-baby method has probability $0.71875 + o(1)$.

1. Introduction

Fix a prime ℓ. The discrete-logarithm problem for a group G of order ℓ is the problem of finding $\log_g h$, given a generator g of G and an element h of G. The notation $\log_g h$ means the unique $s \in \mathbf{Z}/\ell$ such that $h = g^s$, where G is written multiplicatively.

The difficulty of finding discrete logarithms depends on G. For example, if G is the additive group $\mathbf{Z}/\ell$ (encoded as bit strings representing $\{0, 1, \ldots, \ell - 1\}$ in the usual way), then $\log_g h$ is simply h/g, which can be computed in polynomial time using the extended Euclidean algorithm. As a more difficult example, consider the case that $p = 2\ell + 1$ is prime and G is the order-ℓ subgroup of the multiplicative group $\mathbf{F}_p^*$ (again encoded in the usual way); index-calculus attacks then run in time subexponential in p and thus in ℓ. However, if G is the order-ℓ subgroup of $\mathbf{F}_p^*$ where $p - 1$ is a much larger multiple of ℓ, then index-calculus attacks become much slower in terms of ℓ; the standard algorithms are then the baby-step-giant-step method, using at most $(2 + o(1))\sqrt{\ell}$ multiplications in G, and the rho method, which if tweaked carefully uses on average $(\sqrt{\pi/2} + o(1))\sqrt{\ell}$ multiplications in G.

This paper focuses on generic discrete-logarithm algorithms such as the baby-step-giant-step method and the rho method. "Generic" means that these algorithms work for any order-ℓ group G, using oracles to compute $1 \in G$ and to compute $a, b \mapsto ab$ for any $a, b \in G$. See Section 2 for a precise definition.

If G is an elliptic-curve group chosen according to standard criteria then the best discrete-logarithm algorithms available are variants of the baby-step-giant-step method and the rho method, taking advantage of the negligible cost of computing

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inverses in G. There is a standard "inverting" (or "negating") variant of the concept of a generic algorithm, also discussed in Section 2. This paper emphasizes the non-inverting case, but all of the ideas can be adapted to the inverting case.

Measuring algorithm cost. The most fundamental metric for generic discrete-logarithm algorithms, and the metric used throughout this paper, is the probability of discovering a uniform random discrete logarithm within m multiplications. By appropriate integration over m one obtains the average number of multiplications to find a discrete logarithm, the variance, etc. We caution the reader that comparing probabilities of two algorithms for one m can produce different results from comparing averages, maxima, etc.; for example, the rho method is faster than baby-step-giant-step on average but much slower in the worst case.

One can interpret a uniform random discrete logarithm as $\log_g h$ for a uniform random pair (g, h), or as $\log_g h$ for a fixed g and a uniform random h. The following trivial "worst-case-to-average-case reduction" shows that a worst-case discrete logarithm is at most negligibly harder than a uniform random discrete logarithm: one computes $\log_g h$ as $\log_g h' - r$ where $h' = hg^r$ for a uniform random $r \in \mathbf{Z}/\ell$.

There are many reasons that simply counting multiplications, the number m above, does not adequately capture the cost of these algorithms:

- A multiplication count ignores overhead, i.e., costs of computations other than multiplications. For example, the ongoing ECC2K-130 computation uses a very restricted set of Frobenius powers, sacrificing approximately 2% in the number of multiplications, because this reduces the overhead enough to speed up the entire computation.
- A multiplication count ignores issues of memory usage. For some algorithms, such as the baby-step-giant-step method, the memory usage grows with $\sqrt{\ell}$, while others, such as the rho method, use constant (or near-constant) memory.
- A multiplication count is blind to optimizations of the multiplication operation. The question here is not simply how fast multiplication can be, but how multiplication algorithms interact with higher-level choices in these algorithms. For example, Cheon, Hong, and Kim in [8] showed how to look ahead one step in the rho method for $\mathbf{F}_p^*$ and combine two multiplications into one with very little overhead, although memory usage increases.
- A multiplication count ignores issues of parallelization. Pollard's original rho method is difficult to parallelize effectively, but "distinguished point" variants of the rho method are heavily parallelizable with small overhead.
- A multiplication count ignores issues of vectorization. Modern processors can operate on a vector of words in one clock cycle but this requires that the operation is the same across the entire vector. This issue was raised in a recent discussion of whether the negation map on an elliptic curve can actually be used to speed up the rho method, rather than merely to save multiplications; see [4] and [2] for the two sides of the argument.

An improvement in multiplication counts does not necessarily indicate an improvement in more sophisticated cost metrics. It is nevertheless reasonable to begin with an analysis of multiplication counts, as is done in a large fraction of the literature; followup analyses can then ask whether improved multiplication counts are still achievable by algorithms optimized for other cost metrics.

Contents of this paper. Brent and Pollard in [5] identified a source of nonrandomness in the rho method, and quantified the loss of success probability produced by this nonrandomness, under plausible heuristic assumptions. The Brent–Pollard nonrandomness (with various simplifications and in various special cases) has been stated by many authors as the main deficiency in the rho method, and the rho method has been the workhorse of large-scale discrete-logarithm computations. There appears to be a widespread belief that, except for the Brent–Pollard nonrandomness, the rho method is the best conceivable generic discrete-logarithm algorithm. Of course, the rho method can take more than $2\sqrt{\ell}$ multiplications in the worst case while the baby-step-giant-step method is guaranteed to finish within $2\sqrt{\ell}$ multiplications, but the rho method is believed to be the best way to spend a smaller number of multiplications.

This paper shows that there are actually at least two more steps separating the rho method from optimality. First, the rho method is actually less random and less successful than the Brent–Pollard prediction, because the rho method suffers from a tower of what we call "local anti-collisions"; Brent and Pollard account only for "degree-1 local anti-collisions". Second, and more importantly, the rho method would not be optimal even if it were perfectly random, because it continues to suffer from what we call "global anti-collisions". We introduce a new "two grumpy giants and a baby" algorithm that avoids many of these global anti-collisions.

This new algorithm, like the original baby-step-giant-step algorithm, has low overhead but high memory. We have not found a low-memory variant. This means that, for the moment, the algorithm is useful only for discrete-logarithm problems small enough to fit into fast memory. The algorithm nevertheless challenges the idea that the rho method is optimal for larger problems. The same approach might also be useful for "implicit" discrete-logarithm problems in which rho-type iteration is inapplicable, such as stage 2 of the $p-1$ factorization method, but those problems involve many overheads not considered in this paper.

Section 2 describes the general concept of anti-collisions. Section 3 reviews the Brent–Pollard nonrandomness. Section 4 discusses higher-degree anti-collisions in the rho method. Section 5 reports computations of optimal discrete-logarithm algorithms for small ℓ. Section 6 presents our new algorithm.

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2. Anti-collisions

This section introduces the concept of anti-collisions in generic discrete-logarithm algorithms. This section begins by reviewing the standard way to define such algorithms; readers familiar with the definition should still skim it to see our notation.

Generic discrete-logarithm algorithms. The standard way to formalize the idea that a generic algorithm works for any order-ℓ group G is to give the algorithm access to an oracle that computes $1 \in G$ and an oracle that computes the function $a, b \mapsto ab$ from $G \times G$ to G. The elements of G are encoded as a size-ℓ set $\overline{G}$ of strings.

An m-multiplication generic algorithm is one that calls the $a, b \mapsto ab$ oracle m times. The algorithm obtains 1 for free, and has g and h as inputs, so overall it sees $m + 3$ group elements. We write $w_0 = 1$, $w_1 = g$, $w_2 = h$, and w_i for $i \geq 3$ as the $(i-2)$nd output of the $a, b \mapsto ab$ oracle: in other words, $w_i = w_j w_k$ for some $j, k \in \{0, 1, \ldots, i-1\}$ computed by the algorithm as functions of $w_0, w_1, \ldots, w_{i-1}$.

These functions can also flip coins (i.e., take as an additional input a sequence $b_0, b_1, \ldots$ of uniform random bits that are independent of each other, of g, of h, etc.), but cannot make oracle calls.

The standard way to formalize the idea that a generic algorithm does not take advantage of the structure of G is to hide this structure by randomizing it. For example, one can take G as the additive group $\mathbf{Z}/\ell$, and take $\overline{G}$ as the usual binary representation of $\{0, 1, \ldots, \ell - 1\}$, but choose a uniform random injection from G to $\overline{G}$ rather than the usual encoding. One defines the *generic* success probability of a generic algorithm by averaging not only over $\log_g h$ but also over the choices of this injection.

To allow inverting algorithms one also allows free access to an oracle that computes $a \mapsto 1/a$. Equivalently, one allows the algorithm to compute w_i as either $w_j w_k$ or w_j/w_k, and one also provides $1/w_i$. Of course, one can simulate this inversion oracle using approximately $\lg \ell$ calls to the multiplication oracle, since $1/a = a^{\ell-1}$; an algorithm that uses only a small number of inversions can thus be simulated at negligible cost without inversions.

Slopes. Each w_i can be written as $h^{x_i} g^{y_i}$ for a pair $(x_i, y_i) \in (\mathbf{Z}/\ell)^2$ trivially computable by the algorithm. Specifically, $w_0 = 1 = h^{x_0} g^{y_0}$ where $(x_0, y_0) = (0, 0)$; $w_1 = g = h^{x_1} g^{y_1}$ where $(x_1, y_1) = (0, 1)$; $w_2 = h = h^{x_2} g^{y_2}$ where $(x_2, y_2) = (1, 0)$; if w_i is computed as $w_j w_k$ then $w_i = h^{x_i} g^{y_i}$ where $(x_i, y_i) = (x_j, y_j) + (x_k, y_k)$; and if an inverting algorithm computes w_i as w_j/w_k then $w_i = h^{x_i} g^{y_i}$ where $(x_i, y_i) = (x_j, y_j) - (x_k, y_k)$.

Normally these algorithms find $\log_g h$ by finding collisions in the map $(x, y) \mapsto h^x g^y$ from $(\mathbf{Z}/\ell)^2$ to G. A collision $h^{x_i} g^{y_i} = h^{x_j} g^{y_j}$ with $(x_i, y_i) \neq (x_j, y_j)$ must have $x_i \neq x_j$ (otherwise $g^{y_i} = g^{y_j}$ so $y_i = y_j$ since g generates G), so the negative of the slope $(y_j - y_i)/(x_j - x_i)$ is exactly $\log_g h$. The discrete logarithms found by $w_0, w_1, \ldots, w_{m+2}$ are thus exactly the negatives of the $(m+3)(m+2)/2$ slopes (excluding any infinite slopes) between the $m+3$ points $(x_0, y_0), \ldots, (x_{m+2}, y_{m+2})$ in $(\mathbf{Z}/\ell)^2$. The number of discrete logarithms found in this way is the number d of distinct non-infinite slopes. The generic chance of encountering such a collision is exactly d/ℓ.

In the remaining cases, occurring with probability $1 - d/\ell$, these algorithms simply guess $\log_g h$. The success chance of this guess is 0 if the guess matches one of the negated slopes discussed above; otherwise the conditional success chance of this guess is $1/(\ell - d)$, so the success chance of this guess is $1/\ell$. The overall generic success chance of the algorithm is thus between d/ℓ and $(d+1)/\ell$, depending on the strategy for this final guess. In the extreme case $d = \ell$ this guess does not exist and the generic success chance is 1.

(Similar comments apply to inverting algorithms, but the bound on d is doubled, because there are twice as many opportunities to find $-\log_g h$. Specifically, comparing w_j to w_i finds the slope $(y_j - y_i)/(x_j - x_i)$, while comparing w_j to $1/w_i$ finds $(y_j + y_i)/(x_j + x_i)$.)

This model for generic discrete-logarithm algorithms was introduced by Shoup in [19], along with the bound $O(m^2/\ell)$ on the generic success probability of m-multiplication algorithms. Nechaev in [13] three years earlier had proven the collision-probability bound $O(m^2/\ell)$ in a weaker model, where algorithms are permitted only to remotely manipulate group elements without inspecting strings representing the group elements. Nechaev's model is equivalent to Shoup's model when

one measures algorithm cost as the number of multiplications, but is more restrictive than Shoup's model in more sophisticated cost metrics; for example, Nechaev's model is unable to express the rho algorithm.

Chateauneuf, Ling, and Stinson in [7] introduced the idea of counting distinct slopes. They pointed out that the success probability of the baby-step-giant-step method is a factor $2 + o(1)$ away from the obvious quantification of the Nechaev–Shoup bound: m multiplications allow only $m/2$ baby steps and $m/2$ giant steps, producing $m^2/4$ slopes (all distinct if $m^2/4 \leq \ell$), while one can imagine $m + 3$ points in $(\mathbf{Z}/\ell)^2$ potentially having as many as $(m + 3)(m + 2)/2 > m^2/2$ distinct slopes.

Computer searches reported in [7, Section 3] found for each $\ell < 100$ a set of only marginally more than $\sqrt{2\ell}$ points with slopes covering $\mathbf{Z}/\ell$. However, these sets of points do not form addition chains, and as far as we can tell the shortest addition chains for all of the constructions in [7] are worse than the baby-step-giant-step method in the number of multiplications used. The cost model used in [7] allows $a, b \mapsto a^s b^t$ as a single oracle call for any (s, t); we view that cost model as excessively simplified, and are skeptical that algorithms optimized for that cost model will be of any use in practice.

Anti-collisions. We use the word "anti-collision" to refer to an appearance of a useless slope — a slope that cannot create a new collision because the same slope has appeared before. Formally, an anti-collision is a pair (i, j) with $i > j$ such that either

- $x_i = x_j$ or
- $(y_j - y_i)/(x_j - x_i)$ equals $(y_{j'} - y_{i'})/(x_{j'} - x_{i'})$ for some pair (i', j') lexicographically smaller than (i, j) with $i' > j'$.

The number of anti-collisions is exactly the gap $(m + 3)(m + 2)/2 - d$, where as above d is the number of distinct non-infinite slopes. Our objective in this paper is to understand *why* anti-collisions occur in addition chains in $(\mathbf{Z}/\ell)^2$, and how these anti-collisions can be avoided.

In Section 3 we review a standard heuristic by Brent and Pollard that can be viewed as identifying some anti-collisions in the rho method, making the rho method somewhat less effective than a truly random walk would be. In Section 4 we identify a larger set of anti-collisions in the rho method, making the rho method even less effective than predicted by Brent and Pollard. This difference is most noticeable for rho walks that use a very small number of steps, such as hardware-optimized walks or typical walks on equivalence classes modulo Frobenius on Koblitz curves.

It should be obvious that even a truly random walk produces a large number of anti-collisions when m grows to the scale of $\sqrt{\ell}$. In Section 6 we show that at least a constant fraction of these anti-collisions can be eliminated: we construct an explicit and efficient addition chain with significantly fewer anti-collisions, and thus significantly higher success probability, than a truly random walk.

3. Review of the Brent–Pollard nonrandomness

This section reviews the nonrandomness that Brent and Pollard pointed out in the rho method. The literature contains three formulas for this nonrandomness, in three different levels of generality, backed by two different heuristic arguments. As

discussed in Section 4, these heuristics account for "degree-1 local anti-collisions" but do not account for "higher-degree local anti-collisions".

The rho method. The rho method precomputes r distinct "steps" $s_1, s_2, \ldots, s_r \in G - \{1\}$ (as some initial w's), and then moves from w_i to $w_{i+1} = w_i s_j$, where j is a function of w_i. Write p_j for the probability that step s_j is used.

We suppress standard details of efficient parallelization and collision detection here, since our emphasis is on the success probability achieved after m multiplications. Inserting each new group element into an appropriate data structure will immediately recognize the first collision without consuming any multiplications.

The $\sqrt{V}$ formula. Brent and Pollard in [5, Section 2] introduced the following heuristic argument, concluding that if the values $w_0, \ldots, w_m$ are distinct then w_{m+1} collides with one of those values with probability approximately mV/ℓ, where V is defined below. This implies that the total chance of a collision within m multiplications (i.e., within $w_0, \ldots, w_{m+2}$) is approximately $1 - (1 - V/\ell)^{m^2/2}$, which in turn implies that the average number of multiplications for a collision is approximately $\sqrt{\pi/2}\sqrt{\ell}/\sqrt{V}$. For comparison, a truly random walk would have $V = 1$.

This argument applies to a more general form of the rho method, in which some function F is applied to w_i to produce w_{i+1}. The first collision might be unlucky enough to involve w_0, but otherwise it has the form $w_{i+1} = w_{j+1}$ with $w_i \neq w_j$, revealing a collision $F(w_i) = F(w_j)$ in the function F. Applications vary in how they construct F and in the use that they make of a collision.

Assume, heuristically, that the probability of w_i matching any particular value y is proportional to the number of preimages of y, i.e., that $\Pr[w_i = y] = \#F^{-1}(y)/\ell$ where $F^{-1}(y)$ means $\{x : F(x) = y\}$. This heuristic is obviously wrong for w_0, but this is a minor error in context; the heuristic seems plausible for $w_1, \ldots, w_m$, which are each generated as outputs of F.

Assume that $w_0, \ldots, w_m$ are distinct. Define X as the set of preimages of $w_1, \ldots, w_m$, i.e., the disjoint union of $F^{-1}(w_1), \ldots, F^{-1}(w_m)$. Then the expected size of X is

$$\sum_x \Pr[x \in X] = \sum_x \sum_i \Pr[F(x) = w_i] = \sum_x \sum_i \sum_y \Pr[F(x) = y \text{ and } w_i = y].$$

Assume, heuristically, that $F(x) = y$ and $w_i = y$ are independent events. Then the expected size of X is $\sum_i \sum_y \sum_x \Pr[F(x) = y] \Pr[w_i = y] = \sum_i \sum_y \#F^{-1}(y)^2/\ell = m \sum_y \#F^{-1}(y)^2/\ell$.

Define V as the variance over y of $\#F^{-1}(y)$. The average over y of $\#F^{-1}(y)$ is 1, so $V = (\sum_y \#F^{-1}(y)^2/\ell) - 1$, so the expected size of X is $mV + m$. There are m known elements $w_0, \ldots, w_{m-1}$ of X; the expected number of elements of X other than $w_0, \ldots, w_{m-1}$ is mV. By hypothesis w_m is none of $w_0, \ldots, w_{m-1}$; if w_m were uniformly distributed subject to this constraint then it would have probability $mV/(\ell - m) \approx mV/\ell$ of being in X and thus leading to a collision in the next step.

The $\sqrt{1 - \sum_i p_i^2}$ formula. As part of [1] we introduced the following streamlined heuristic argument, concluding that the collision probability for w_{m+1} is approximately $m(1 - \sum_i p_i^2)/\ell$. This implies that the average number of multiplications for a collision is approximately $\sqrt{\pi/2}\sqrt{\ell}/\sqrt{1 - \sum_i p_i^2}$.

Fix a group element v, and let w and w' be two independent uniform random elements. Consider the event that w and w' both map to v but $w \neq w'$. This

event occurs if there are distinct i, j such that the following three conditions hold simultaneously:

- $v = s_i w = s_j w'$;
- s_i is chosen for w;
- s_j is chosen for w'.

These conditions have probability $1/\ell^2$, p_i, and p_j respectively. Summing over all (i, j) gives the overall probability $\left(\sum_{i \neq j} p_i p_j\right)/\ell^2 = \left(\sum_{i,j} p_i p_j - \sum_i p_i^2\right)/\ell^2 = \left(1 - \sum_i p_i^2\right)/\ell^2$. This means that the probability of an immediate collision from w and w' is $\left(1 - \sum_i p_i^2\right)/\ell$, where we added over the ℓ choices of v.

After $m + 3$ group elements one has approximately $m^2/2$ potentially colliding pairs. If the inputs to the iteration function were independent uniformly distributed random points then the probability of success would be $1 - \left(1 - \left(1 - \sum_i p_i^2\right)/\ell\right)^{m^2/2}$ and the average number of iterations before a collision would be approximately $\sqrt{\pi/2}\sqrt{\ell}/\sqrt{1 - \sum_i p_i^2}$. The inputs to the iteration function in Pollard's rho method are not actually independent, but this has no obvious effect on the average number of iterations.

Relating the two formulas. We originally obtained the formula $\sqrt{1 - \sum_i p_i^2}$ by specializing and simplifying the Brent–Pollard $\sqrt{V}$ formula as follows.

The potential preimages of y are $y/s_1, y/s_2, \ldots, y/s_r$, which are actual preimages with probabilities $p_1, p_2, \ldots, p_r$ respectively. A subset I of $\{1, 2, \ldots, r\}$ matches the set of indices of preimages with probability $(\prod_{i \in S} p_i)(\prod_{i \notin S}(1 - p_i))$, so the average of $\#F^{-1}(y)^2$ is $\sum_I \#I^2 (\prod_{i \in S} p_i)(\prod_{i \notin S}(1 - p_i))$. It is easy to see that most monomials (e.g., $p_1 p_2 p_3$) have coefficient 0 in this sum; the only exceptions are linear monomials p_i, which have coefficient 1, and quadratic monomials $p_i p_j$ with $i < j$, which have coefficient 2. The sum therefore equals $\sum_i p_i + 2 \sum_{i,j:i<j} p_i p_j = \sum_i p_i + (\sum_i p_i)^2 - \sum_i p_i^2 = 2 - \sum_i p_i^2$. Hence $V = 1 - \sum_i p_i^2$.

The $\sqrt{1 - 1/r}$ formula. In traditional "adding walks" (credited to Lenstra in [16, page 66]; see also [17, page 295] and [20]), each p_i is $1/r$, and $\sqrt{1 - \sum_i p_i^2}$ is $\sqrt{1 - 1/r}$. This $\sqrt{1 - 1/r}$ formula first appeared in [20], with credit to the subsequent paper [3] by Blackburn and Murphy. The heuristic argument in [3] is the same as the Brent–Pollard argument.

Case study: Koblitz curves. The $\sqrt{1 - \sum_i p_i^2}$ formula was first used to optimize walks on Koblitz curves. These walks map a curve point W to $W + \varphi^i(W)$, where φ is the Frobenius map and i is chosen as a function of the Hamming weight of the normal-basis representation of the x-coordinate of W. The Hamming weight is not uniformly distributed, and any reasonable function of the Hamming weight is also not uniformly distributed, so the $\sqrt{1 - 1/r}$ formula does not apply. Note that these are "multiplying walks" rather than "adding walks" (if $W = x_i H + y_i G$ then $W + \varphi^i(W) = s_i x_i H + s_i y_i G$ for certain constants $s_i \in (\mathbf{Z}/\ell)^*$), but the heuristics in this section are trivially adapted to this setting.

As a concrete example we repeat from [1] the analysis of our ongoing attack on ECC2K-130. All Hamming weights of x-coordinates of group elements are even, and experiments show that the distribution of even-weighted words of length 131 is close to the distribution of x-coordinates of group elements. Any iteration function defined in this way therefore inevitably introduces an extra factor to the running

time of $1/\sqrt{1-\sum_i \binom{131}{2i}^2/2^{260}} \approx 1.053211$, even if all 66 weights use different scalars s_i. We extract just 3 bits of weight information, using only 8 different values for the scalars, to reduce the time per iteration. The values are determined by $\mathrm{HW}(x_{P_i})/2 \bmod 8$; the distribution of $\sum_i \binom{131}{16i+2j}$ for $0 \le j \le 7$ gives probabilities

$$0.1414, 0.1443, 0.1359, 0.1212, 0.1086, 0.1057, 0.1141, 0.1288,$$

giving a total increase of the number of iterations by a factor of 1.069993.

4. Higher-degree local anti-collisions

Consider the rho method using r "steps" $s_1, s_2, \ldots, s_r \in G$, as in the previous section. The method multiplies w_i by one of these steps to obtain w_{i+1}, multiplies w_{i+1} by one of these steps to obtain w_{i+2}, etc.

Assume that the step w_{i+1}/w_i is different from the step w_{i+2}/w_{i+1}, but that w_{i+1}/w_i is the same as an earlier step w_{j+2}/w_{j+1}, and that w_{i+2}/w_{i+1} is the same as the step w_{j+1}/w_j. There are anti-collisions $(i+1, j+2)$ and $(i+2, j+1)$, exactly the phenomenon discussed in the previous section: for example, w_{i+1} cannot equal w_{j+2} unless w_i equals w_{j+1}. There is, however, also a local anti-collision $(i+2, j+2)$ not discussed in the previous section: w_{i+2} cannot equal w_{j+2} unless w_i equals w_j. The point is that the ratio w_{i+2}/w_i is a product of two steps, and the ratio w_{i+2}/w_i is a product of the same two steps in the opposite order.

We compute the heuristic impact of these "degree-2 local anti-collisions", together with the degree-1 local anti-collisions of Section 3, as follows. Assume for simplicity that $1, s_1, s_2, \ldots, s_r, s_1^2, s_1s_2, \ldots, s_1s_r, s_2^2, \ldots, s_2s_r, \ldots, s_r^2$ are distinct. Write $F(w)$ for the group element that w maps to. Fix a group element v, and consider the event that two independent uniform random group elements w, w' have $F(F(w)) = v = F(F(w'))$ with no collisions among $w, w', F(w), F(w')$. This event occurs if there are i, i', j, j' with $s_j \ne s_{j'}$ and $s_j s_i \ne s_{j'} s_{i'}$ such that the following conditions hold simultaneously:

- $v = s_j s_i w = s_{j'} s_{i'} w'$;
- $F(w) = s_i w$;
- $F(s_i w) = s_j s_i w$;
- $F(w') = s_{i'} w'$;
- $F(s_{i'} w') = s_{j'} s_{i'} w'$.

These conditions have probability $1/\ell^2$, p_i, p_j, $p_{i'}$, and $p_{j'}$ respectively. Given the first condition, the remaining conditions are independent of each other, since $w = v/(s_j s_i)$, $s_i w = v/s_j$, $w' = v/(s_{j'} s_{i'})$, and $s_{i'} w' = v/s_{j'}$ are distinct. This event thus has probability $\sum p_i p_j p_{i'} p_{j'}/\ell^2$ where the sum is over all i, j, i', j' with $s_j \ne s_{j'}$ and $s_j s_i \ne s_{j'} s_{i'}$. The complement of the sum is over all i, j, i', j' with $s_j = s_{j'}$ or $s_j s_i = s_{j'} s_{i'}$, i.e., with $j = j'$ or with $i' = j \ne j' = i$. The complement is thus $\sum_j p_j^2 + \sum_{i,j:i\ne j} p_i^2 p_j^2 = \sum_j p_j^2 + (\sum_j p_j^2)^2 - \sum_j p_j^4$, and the original sum is $1 - \sum_j p_j^2 - (\sum_j p_j^2)^2 + \sum_j p_j^4$. Adding over all v gives probability $(1 - \sum_j p_j^2 - (\sum_j p_j^2)^2 + \sum_j p_j^4)/\ell$ of this type of two-step collision between w and w'.

For example, if $p_i = 1/r$ for all i, then the degree-1-and-2 nonrandomness factor is $1/\sqrt{1 - 1/r - 1/r^2 + 1/r^3}$, whereas the Brent–Pollard (degree-1) nonrandomness factor is $1/\sqrt{1 - 1/r}$. These factors are noticeably different if r is small.

Beyond degree 2. More generally, a "degree-k local anti-collision" $(i+k, j+k)$ occurs when the product of k successive steps w_{i+1}/w_i, w_{i+2}/w_{i+1}, etc. matches the product of k successive steps w_{j+1}/w_j, w_{j+2}/w_{j+1}, etc., without a lower-degree local anti-collision occurring. A "degree-(k,k') local anti-collision" $(i+k, j+k')$ is defined similarly.

Given the vector $(s_1, s_2, \ldots, s_r)$, one can straightforwardly compute the overall heuristic effect of local anti-collisions of degree at most k, by summing the products $p_{i_1} \cdots p_{i_k} p_{i'_1} \cdots p_{i'_k}$ for which 1, s_{i_1}, $s_{i'_1}$, $s_{i_1} s_{i_2}$, $s_{i'_1} s_{i'_2}$, etc. are distinct. Experiments indicate that the largest contribution is usually from the smallest degrees.

We emphasize that the results depend on the vector $(s_1, s_2, \ldots, s_r)$, because generic commutative-group equations such as $s_1 s_2 = s_2 s_1$ are not the only multiplicative dependencies among $s_1, s_2, \ldots, s_r$. One can check that $s_1, s_2, \ldots, s_r$ have no non-generic multiplicative dependencies of small degree (and modify them to avoid such dependencies), but they always have medium-degree non-generic multiplicative dependencies, including mixed-degree non-generic multiplicative dependencies.

If $s_1, s_2, \ldots, s_r$ have only generic dependencies of degree at most k then the sum described above is expressible as a polynomial in the easily computed quantities $X_2 = \sum_j p_j^2$, $X_4 = \sum_j p_j^4$, etc., by a simple inclusion-exclusion argument. For example, the degree-1 nonrandomness factor is $1/\sqrt{1 - X_2}$, as in Section 3; the degree-$\le$2 nonrandomness factor is $1/\sqrt{1 - X_2 - X_2^2 + X_4}$, as explained above; the degree-$\le$3 nonrandomness factor is $1/\sqrt{1 - X_2 - X_2^2 - 3X_2^3 + X_4 + 7X_2X_4 - 4X_6}$; etc. In the uniform case these factors are $1/\sqrt{1 - 1/r}$, $1/\sqrt{1 - 1/r - 1/r^2 + 1/r^3}$, $1/\sqrt{1 - 1/r - 1/r^2 - 2/r^3 + 7/r^4 - 4/r^5}$, etc.

Case study: $r = 6$. Hildebrand showed in [11] that almost every r-adding walk (with $p_j = 1/r$) reaches a nearly uniform distribution in $\mathbf{Z}/\ell$ within $O(\ell^{2/(r-1)})$ steps; in particular, within $o(\sqrt{\ell})$ steps for $r \ge 6$. Implementors optimizing Pollard's rho method for hardware often want r to be as small as possible (to minimize the storage required for precomputed steps and the cost of accessing that storage), and in light of Hildebrand's result can reasonably choose $r = 6$. This raises the question of how random a 6-adding walk is; perhaps it is better to take a larger value of r, increasing overhead but reducing nonrandomness.

For $r = 6$, with $p_1 = p_2 = p_3 = p_4 = p_5 = p_6 = 1/6$, the nonrandomness factors are approximately 1.095445 for degree 1, 1.110984 for degree ≤ 2, 1.117208 for degree ≤ 3, 1.120473 for degree ≤ 4, etc. Evidently the Brent–Pollard heuristic captures *most* of the impact of local anti-collisions for $r = 6$, but not all of the impact.

We tried 10000000 experiments for $\ell = 1009$. Each experiment generated 6 uniform random steps $s_1, s_2, \ldots, s_6$ (without enforcing distinctness, and without any constraints on higher-degree multiplicative dependencies), carried out a random walk using $s_1, s_2, \ldots, s_6$ with equal probability, and stopped at the first collision. We reused each $s_1, s_2, \ldots, s_6$ for a batch of 1000 experiments. The average walk length was approximately 1.1513 times $\sqrt{\pi/2}\sqrt{\ell}$; note that this does not count the multiplications used to generate $s_1, s_2, \ldots, s_6$. We then tried 10000000 experiments for $\ell = 8191$, obtaining 1.1488 times $\sqrt{\pi/2}\sqrt{\ell}$; for $\ell = 65537$, obtaining 1.1429; and for $\ell = 524287$, obtaining 1.1377. Our heuristics predict that these numbers will converge to something above 1.12 as $\ell \to \infty$, rather than 1.095445, but obviously

the experimental data is too limited to provide any confidence. Note that for small ℓ there are more low-degree dependencies among the steps s_i, so it is not a surprise that smaller values of ℓ have larger nonrandomness factors. We are carrying out larger experiments.

Case study: Koblitz curves, revisited. Consider again the ECC2K-130 walk introduced in [1]. Here $\ell = 68056473384187692693232012949340998512 9$.

For $0 \le j \le 7$ define φ as the Frobenius map on the ECC2K-130 curve, and define $s_j \in \mathbf{Z}/\ell$ as $1 + 196511074115861092422032515080945363956^{j+3}$. This walk moves from P to $P + \varphi^{j+3}(P) = s_j P$ if the Hamming weight of the x-coordinate of P is congruent to $2j$ modulo 16; this occurs with probability (almost exactly) $p_j = \sum_i \binom{131}{16i+2j}/2^{130}$.

The only small-degree multiplicative dependencies among $s_0, \ldots, s_7$ are generic commutative-group equations such as $s_1 s_2 = s_2 s_1$. We already reported this in [1, Section 2] to explain why the walk is highly unlikely to enter a short cycle. We point out here that this has a larger effect, namely minimizing small-degree anti-collisions. We now analyze the impact of the small-degree anti-collisions that remain, those that arise from the generic commutative-group equations.

For degree 1 the nonrandomness factor is $1/\sqrt{1 - X_2} \approx 1.069993$. For degree ≤ 2 the nonrandomness factor is $1/\sqrt{1 - X_2 - X_2^2 + X_4} \approx 1.078620$. For degree ≤ 3 the nonrandomness factor is $1/\sqrt{1 - X_2 - X_2^2 - 3X_2^3 + X_4 + 7X_2X_4 - 4X_6} \approx 1.081370$. For degree ≤ 4 the nonrandomness factor is ≈ 1.082550.

Case study: Mixed walks. The same type of analysis should also apply to "mixed walks" combining non-commuting steps such as $w \mapsto ws_1$, $w \mapsto ws_2$, and $w \mapsto w^2$. However, we have not yet carried out experiments.

Optimizing asymptotics. It is frequently stated that the rho method, like a truly random walk, finishes in $(\sqrt{\pi/2} + o(1))\sqrt{\ell}$ multiplications on average.

However, the experimental results by Sattler and Schnorr [16, page 76] and by Teske [20] showed clearly that $\sqrt{\pi/2} + o(1)$ is not achieved by small values of r, and in particular by Pollard's original rho method. The Brent–Pollard nonrandomness, and in particular the $\sqrt{1 - 1/r}$ formula, indicates that $\sqrt{\pi/2} + o(1)$ is not achieved by any bounded r; one must have $1/r \in o(1)$, i.e., $r \to \infty$ as $\ell \to \infty$. On the other hand, if r grows too quickly then the cost of setting up r steps is nonnegligible.

This analysis does not contradict $\sqrt{\pi/2} + o(1)$. However, it does indicate that some care is required in the algorithm details, and that $\sqrt{\pi/2} + o(1)$ can be replaced by $\sqrt{\pi/2} + O(\ell^{-1/4})$ but not by $\sqrt{\pi/2} + o(\ell^{-1/4})$.

To optimize the $o(1)$ one might try choosing steps that are particularly easy to compute. For example, one might take $s_3 = s_1 s_2$, $s_4 = s_2 s_3$, etc., where s_1, s_2 are random. We point out, however, that such choices are particularly prone to higher-degree anti-collisions. We recommend taking into account not just the number of steps and the number of multiplications required to precompute those steps, but also the impact of higher-degree anti-collisions.

5. Searching for better chains for small primes

If ℓ is small then by simply enumerating addition chains one can find generic discrete-logarithm algorithms that use fewer multiplications than the rho method.

This section reports, for each small prime ℓ, the results of two different computer searches. One search greedily obtained as many slopes as it could after each multiplication, deferring anti-collisions as long as possible. The other search minimized the number of multiplications required to find an average slope. Chains found by such searches are directly usable in discrete-logarithm computations for these values of ℓ; perhaps they also provide some indication of what one can hope to achieve for much larger values of ℓ. These searches also show that merely counting the size of a slope cover, as in [7, Section 3], underestimates the cost of discrete-logarithm algorithms, although one can hope that the gap becomes negligible as ℓ increases.

A continuing theme in this section is that the obvious quantification of the Nechaev–Shoup bound is not tight. The bound says that an m-addition chain has $\leq(m+3)(m+2)/2$ slopes; but there is actually a gap, increasing with m, between $(m+3)(m+2)/2$ and the maximum number of slopes in an m-addition chain. This section explains part of this gap by identifying two types of anti-collisions that addition chains cannot avoid and stating an improved bound that accounts for these anti-collisions. However, the improved bound is still not tight for most of these values of ℓ, and for long chains the improved bound is only negligibly stronger than the Nechaev–Shoup bound.

Greedy slopes. Define d_i as the number of distinct finite slopes among the points $(x_0, y_0), (x_1, y_1), (x_2, y_2), \ldots, (x_i, y_i)$ in $(\mathbf{Z}/\ell)^2$. For example, the chain

$$(0,0), (0,1), (1,0), (0,2), (1,2), (1,4)$$

in $(\mathbf{Z}/7)^2$ has $(d_0, d_1, d_2, d_3, d_4, d_5) = (0, 0, 2, 3, 5, 7)$: there are 2 distinct finite slopes among $(0,0), (0,1), (1,0)$; 3 distinct finite slopes among $(0,0), (0,1), (1,0), (0,2)$; 5 distinct finite slopes among $(0,0), (0,1), (1,0), (0,2), (1,2)$; and 7 distinct finite slopes among $(0,0), (0,1), (1,0), (0,2), (1,2), (1,4)$.

For each prime $\ell < 128$ we computed the lexicographically maximum sequence $(d_0, d_1, \ldots)$ for all infinite addition chains starting $(0,0), (0,1), (1,0)$ in $(\mathbf{Z}/\ell)^2$. These maxima, truncated to the first occurrence of ℓ, are displayed in Table 5.1. For example, Table 5.1 lists $(0, 0, 2, 3, 5, 7)$ for $\ell = 7$, indicating that the lexicographic maximum is $(0, 0, 2, 3, 5, 7, 7, 7, 7, 7, \ldots)$: one always has $d_0 = 0$, $d_1 = 0$, and $d_2 = 2$; the maximum possible d_3 is 3; given $d_3 = 3$, the maximum possible d_4 is 5; given $d_3 = 3$ and $d_4 = 5$, the maximum possible d_5 is 7.

This computation was not quite instantaneous, because it naturally ended up computing all finite chains achieving the truncated maximum (and, along the way, all chains achieving every prefix of the truncated maximum). There are, e.g., 5420 length-21 chains that match the $(d_0, d_1, \ldots)$ shown in Table 5.1 for $\ell = 109$.

Minimal weight. We also computed ℓ-slope addition chains of minimal weight for each prime $\ell < 48$. Here "weight" means $\sum_{i \geq 1} i(d_i - d_{i-1})$. Dividing this weight by ℓ produces the average, over all $s \in \mathbf{Z}/\ell$, of the number of multiplications (plus 2 to account for the inputs g and h) used to find slope s. It might make more sense to compute $(\ell - 1)$-slope addition chains of minimal weight, since a generic discrete-logarithm algorithm that finds $\ell - 1$ slopes also recognizes the remaining slope by exclusion, but the gap becomes negligible as ℓ increases.

Lexicographically maximizing $(d_0, d_1, \ldots)$, as in Table 5.1, does not always produce minimal-weight ℓ-slope addition chains. For example, the chain

$$(0,0), (0,1), (1,0), (0,2), (0,3), (1,3), (1,6), (2,12), (2,14), (2,16), (3,17), (4,28)$$

ℓ	weight	$d_0\, d_1\, \ldots$
2	4	0 0 2
3	7	0 0 2 3
5	15	0 0 2 3 5
7	25	0 0 2 3 5 7
11	50	0 0 2 3 5 7 10 11
13	64	0 0 2 3 5 7 10 13
17	96	0 0 2 3 5 7 10 14 16 17
19	113	0 0 2 3 5 7 10 14 17 19
23	148	0 0 2 3 5 7 10 14 19 22 23
29	211	0 0 2 3 5 7 10 14 19 23 26 28 29
31	230	0 0 2 3 5 7 10 14 19 23 28 31
37	300	0 0 2 3 5 7 10 14 19 23 29 33 36 37
41	347	0 0 2 3 5 7 10 14 19 24 29 34 39 41
43	375	0 0 2 3 5 7 10 14 19 24 29 34 38 42 43
47	425	0 0 2 3 5 7 10 14 19 24 30 35 40 44 47
53	510	0 0 2 3 5 7 10 14 19 24 30 36 41 45 50 52 53
59	596	0 0 2 3 5 7 10 14 19 24 30 36 42 48 52 57 58 59
61	631	0 0 2 3 5 7 10 14 19 24 30 35 42 48 52 56 59 61
67	727	0 0 2 3 5 7 10 14 19 24 30 36 41 47 53 59 63 66 67
71	788	0 0 2 3 5 7 10 14 19 24 30 36 42 48 54 60 66 70 71
73	815	0 0 2 3 5 7 10 14 19 24 30 36 43 50 56 62 67 71 73
79	919	0 0 2 3 5 7 10 14 19 24 30 37 43 49 57 64 69 73 76 79
83	978	0 0 2 3 5 7 10 14 19 24 30 37 44 51 59 65 72 77 80 83
89	1081	0 0 2 3 5 7 10 14 19 24 30 37 44 53 60 66 74 80 84 87 89
97	1224	0 0 2 3 5 7 10 14 19 24 30 37 44 51 61 69 78 83 88 92 96 97
101	1307	0 0 2 3 5 7 10 14 19 24 30 37 45 53 60 69 76 82 89 93 97 100 101
103	1351	0 0 2 3 5 7 10 14 19 24 30 37 45 52 60 67 74 83 89 94 98 102 103
107	1422	0 0 2 3 5 7 10 14 19 24 30 37 45 53 61 70 77 84 91 96 100 104 107
109	1466	0 0 2 3 5 7 10 14 19 24 30 37 44 52 60 68 77 84 91 98 102 106 108 109
113	1536	0 0 2 3 5 7 10 14 19 24 30 37 44 52 62 70 78 86 94 99 105 109 113
127	1806	0 0 2 3 5 7 10 14 19 24 30 37 45 53 63 73 84 92 98 105 112 118 122 126 127

TABLE 5.1. For each $\ell < 128$, the lexicographically maximum $(d_0, d_1, \ldots)$. "Weight" means $\sum_{i \geq 1} i(d_i - d_{i-1})$.

for $\ell = 29$ has weight 210 with $(d_0, d_1, \ldots) = (0, 0, 2, 3, 4, 7, 10, 14, 19, 23, 27, 29)$, while chains achieving the lexicographic maximum in Table 5.1 have weight 211. We similarly found weight 299 (compared to 300) for $\ell = 37$, weight 372 (compared to 375) for $\ell = 43$, and weight 423 (compared to 425) for $\ell = 47$. It is not clear whether this gap becomes negligible as ℓ increases.

Some obstructions. We explain here two simple ways that anti-collisions appear in addition chains. Every addition chain produces at least a linear number of anti-collisions that follow these simple patterns.

First, doubling a point (x_j, y_j) produces two anti-collisions: the slopes from $2(x_j, y_j)$ to (x_j, y_j) and to $(0, 0)$ are the same as the slope from (x_j, y_j) to $(0, 0)$. Doubling another point (x_k, y_k) produces three anti-collisions: the slope from $2(x_k, y_k)$ to $2(x_j, y_j)$ is the same as the slope from (x_k, y_k) to (x_j, y_j). A third

doubling produces four anti-collisions, and so on; doubling n points produces a total of $n(n+3)/2$ anti-collisions of this type.

Second, adding (x_i, y_i) to a distinct point (x_j, y_j) produces two anti-collisions: the slopes from $(x_i, y_i)+(x_j, y_j)$ to (x_i, y_i) and to (x_j, y_j) are the same as the slopes from (x_j, y_j) and from (x_i, y_i) to $(0, 0)$. Subsequently adding the same (x_i, y_i) to another point (x_k, y_k) produces three anti-collisions: the slope from $(x_i, y_i)+(x_k, y_k)$ to $(x_i, y_i) + (x_j, y_j)$ is the same as the slope from (x_k, y_k) to (x_j, y_j), exactly as in Section 3.

Applying these principles easily explains the initial pattern $0, 0, 2, 3, 5, 7$ that appears in Table 5.1. The first addition (whether or not a doubling) must produce at least two anti-collisions, and therefore produces at most one new slope to the previous three points; this explains the 3. The second addition also produces at least two anti-collisions, and therefore at most two new slopes to the previous four points; this explains the 5. One might think that the next step is 8, but having only two anti-collisions in each of the first three additions would imply that those three additions include at most one doubling and no other reuse of summands, for a total of at least five summands, while there are only four non-zero summands available for the first three additions.

More generally, a chain of $m \geq 2$ nontrivial additions involves $2m$ inputs selected from $m + 1$ nonzero points, so there must be at least $m - 1$ repetitions of inputs. These repetitions produce at least $m - 2$ occurrences of three anti-collisions (one doubling is free), on top of m occurrences of two anti-collisions and one anti-collision for the infinite slope from $(0, 0)$ to $(0, 1)$, for a total of at least $3m-1$ anti-collisions, and thus a total of at most $(m+3)(m+2)/2-(3m-1) = (m^2-m+8)/2$ slopes. This explains $5, 7, 10, 14, 19$ in Table 5.1 but does not explain 24.

6. Two grumpy giants and a baby

This section presents the algorithm featured in the title of this paper. This algorithm is, as the name suggests, a modification to the standard baby-step-giant-step method. The modification increases the number of different slopes produced within m multiplications, and for a typical range of m increases the number beyond the effectiveness of the rho method.

In the baby-step-giant-step algorithm the baby steps compute $h^{x_i}g^{y_i}$ for $(x_i, y_i) \in (0, 0) + \{0, 1, 2, \ldots, \lceil\sqrt{\ell}\rceil\}(0, 1)$ and the giant steps compute $h^{x_i}g^{y_i}$ for $(x_i, y_i) \in (1, 0) + \{0, 1, 2, \ldots, \lfloor\sqrt{\ell}\rfloor\}(0, \lceil\sqrt{\ell}\rceil)$. The first observation is that the slopes within one type of step are constant; the second observation is that once all steps are done all ℓ slopes appear. Our idea is to make the lines of fixed slope shorter, i.e. introduce more players. Note that introducing a second baby is not useful: lines between the points in $(x, y) + \{0, 1, 2, \ldots, \lceil\sqrt{\ell}\rceil\}(0, 1)$ and $(0, 0) + \{0, 1, 2, \ldots, \lceil\sqrt{\ell}\rceil\}(0, 1)$ repeat each slope $\approx \sqrt{\ell}$ times. We thus need to introduce more giants to make progress.

The two-grumpy-giants-and-a-baby method is parametrized by a positive integer n, normally proportional to $\sqrt{\ell}$; the reader should imagine n being approximately $0.5\sqrt{\ell}$. The number of multiplications in the method is approximately $3n$. Here is the set of points $(x_i, y_i) \in (\mathbf{Z}/\ell)^2$ produced by the method:

Baby: $(0, 0) + \{0, \ldots, n-1\}(0, 1)$
Giant1: $(1, 0) + \{1, \ldots, n\}(0, n)$
Giant2: $(2, 0) - \{1, \ldots, n\}(0, n+1)$

The initial negation $(0, -(n+1))$ for Giant2 has negligible cost, approximately $\lg \ell$ multiplications. Choosing n and $n+1$ for the steps in the y direction for the two giants gives a good coverage of slopes since n and $n+1$ are coprime. The grumpy giants make big steps (on the scale of $\sqrt{\ell}$) and quickly walk in opposite directions away from each other. Luckily they are not minding the baby.

We now analyze the slopes covered by this method. Again it is not interesting to look at the slopes among one type of points. The slope between a point $(0, i)$ in the Baby set and a point $(1, jn)$ in the Giant1 set is $jn - i$; this means that all slopes in $\{1, \ldots, n^2\}$ are covered. The slope between $(0, i)$ in the Baby set and $(2, -j(n+1))$ in the Giant2 set is $(-j(n+1) - i)/2 \in \left\{-n^2 - 2n + 1, \ldots, -n - 1\right\}/2$; there are n^2 distinct slopes here, almost exactly covering $\left\{-n^2 - 2n + 1, \ldots, -n - 1\right\}/2$. The slope between $(1, in)$ in the Giant1 set and $(2, -j(n+1))$ in the Giant2 set is $-j(n+1) - in \in \left\{-2n^2 - n, \ldots, -2n - 1\right\}$; there are another n^2 distinct slopes here, covering about half the elements of $\left\{-2n^2 - n, \ldots, -2n - 1\right\}$.

To summarize, there are three sets of n^2 distinct slopes here, all between $-2n^2 - n + 1$ and n^2. One can hope for a total of $3n^2$ distinct slopes if $\ell > 3n^2 + n$, but this hope runs into two obstacles. The first obstacle is that the "odd" elements of $\left\{-n^2 - 2n + 1, \ldots, -n - 1\right\}$ can bump into the other sets when computing $(2i + 1)/2 = i + (\ell + 1/2)$; but for $\ell \in 4n^2 + O(n)$ this effect loses only $O(n)$ elements. The second obstacle is that any Giant1–Giant2 slopes between $(-n^2 - 2n)/2$ and $(-n - 2)/2$ will bump into $\left\{-n^2 - 2n + 1, \ldots, -n - 1\right\}/2$ for the the "even" elements of $\left\{-n^2 - 2n + 1, \ldots, -n - 1\right\}$. This is approximately the rightmost $1/4$ of the Giant1–Giant2 interval, but only $n^2/8 + O(n)$ of the Giant1–Giant2 slopes are in this interval. Overall there are $23n^2/8 + O(n)$ distinct slopes, i.e., $(0.71875 + o(1))\ell$ distinct slopes.

For comparison, the same $(3 + o(1))n$ multiplications allow the original baby-step-giant-step method to compute $(1.5 + o(1))n$ baby steps and $(1.5 + o(1))n$ giant steps, producing only $(2.25 + o(1))n^2 = (0.5625 + o(1))\ell$ distinct slopes. The same number of multiplications in the rho method (with $1/o(1)$ different steps, simulating a uniform random walk within a factor $1 + o(1)$) produces $(9 + o(1))n^2/2 = (1.125 + o(1))\ell$ random slopes, and thus $(1 - \exp(1.125) + o(1))\ell = (0.6753\ldots + o(1))\ell$ distinct slopes with overwhelming probability. We have performed computer experiments to check each of these numbers.

Weighing the giants. We repeat a warning from Section 1: one algorithm can be better than another after a particular number of multiplications but nevertheless have worse average-case performance.

For example, the baby-step-giant-step method has two standard variants, which we call the baby-steps-then-giant-steps method (introduced by Shanks in [18, pages 419–420]) and the interleaved-baby-step-giant-step method (introduced much later by Pollard in [15, page 439, top]). Both variants (with giant steps chosen to be of size $(1 + o(1))\sqrt{\ell}$) reach 100% success probability using $(2 + o(1))\sqrt{\ell}$ multiplications, while the rho method has a lower success probability for that number of multiplications. Average-case performance tells a quite different story: the baby-steps-then-giant-steps method uses $(1.5 + o(1))\sqrt{\ell}$ multiplications on average; the interleaved-baby-step-giant-step method is better, using $(4/3 + o(1))\sqrt{\ell} = (1.3333\ldots + o(1))\sqrt{\ell}$ multiplications on average; the rho method is best, using $(\sqrt{\pi/2} + o(1))\sqrt{\ell} = (1.2533\ldots + o(1))\sqrt{\ell}$ multiplications on average.

Our analysis above shows that the two-grumpy-giants-and-a-baby method is more effective than the rho method (and the baby-step-giant-step method) as a way to use $(1.5 + o(1))\sqrt{\ell}$ multiplications. One might nevertheless guess that the rho method has better average-case performance; for example, an anonymous referee stated that the new method "presumably has worse average-case running time".

Our computer experiments indicate that the (interleaved-)two-grumpy-giants-and-a-baby method actually has better average-case running time than the rho method. For example, for $\ell = 65537$, we found a chain of weight $20644183 = (1.23046\ldots)\ell^{1.5}$ with the two-grumpy-giants-and-a-baby method. Here we chose $n = 146$, used (suboptimal) binary addition chains for $(0, n)$ and $(0, \ell - n - 1)$, and then cycled between points $(0, i)$ and $(1, in)$ and $(2, -i(n+1))$ until we had ℓ different slopes. For $\ell = 1000003$ we found a chain of weight $1205458963 = (1.20545\ldots)\ell^{1.5}$ in the same way with $n = 558$.

Variants. We have been exploring many variants of this algorithm. We have found experimentally that a 4-giants algorithm (two in one direction, two in the other, with computer-optimized shifts of the initial positions) outperforms this 2-giants algorithm for $m \approx \sqrt{\ell}$. We speculate that gradually increasing the number of giants will produce an algorithm with $(0.5 + o(1))m^2$ distinct slopes, the best possible result (automatically also optimizing the average number of multiplications, the maximum, etc.), but it is not clear how to choose the shift distances properly.

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